

Monte Carlo quadrature

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① Empirical mean

Given X_i ($i=1, 2, \dots, M$)

[sequence of independent random variables
(PDF: π_x)

And, $\begin{pmatrix} x_1 \\ x_2 \\ \vdots \\ x_M \end{pmatrix} = \begin{pmatrix} X_{1(M)} \\ X_{2(M)} \\ \vdots \\ X_{M(M)} \end{pmatrix} \leftarrow$ independent samples

Then, Empirical mean is

$$\bar{x}_M = \frac{1}{M} \sum_{i=1}^M x_i$$

(\bar{x}_M is realisation of random variable $\bar{X}_M = \frac{1}{M} \sum_{i=1}^M X_i$)

This consist Monte Carlo approximation to

$$\int_{\mathbb{R}} x \pi_x dx$$

Also, \bar{x}_M provides an approximation to \bar{x}

as $M \rightarrow \infty$

(cf) which estimator is better?

$$1) \hat{P}_M = \frac{1}{M} \sum_{i=1}^M (X_i - \bar{x}_M)(X_i - \bar{x}_M)^T$$

$$2) \hat{P}_M = \frac{1}{M-1} \sum_{i=1}^M (X_i - \bar{x}_M)(X_i - \bar{x}_M)^T$$

In both cases, the value of \hat{P}_M is the limit of $M \rightarrow \infty$. However, in case 1), $\mathbb{E}[\hat{P}_M] < P$

(P is covariance matrix of X_i)

So, 2) is better.

② Convergence of sequences of random variables

• X_m converges like

$$\mathbb{P}\left(\lim_{M \rightarrow \infty} X_m = X\right) = 1$$

\nwarrow random variable

• Convergence in probability is defined like

$$\forall \varepsilon > 0, \lim_{M \rightarrow \infty} \mathbb{P}(|X_m - x| > \varepsilon) = 0$$

• Weak convergence (or distributions) is defined like

$\forall f, f$ is bounded and continuous

$$\lim_{M \rightarrow \infty} \mathbb{E}[f(X_m)] = \mathbb{E}[f(X)]$$

③ Central limit theorem

Given, random variable \bar{X}_M , PDF π_x , mean \bar{x}

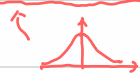
finite variance σ^2 , then random variable Y_M is defined as

$$Y_M = \sqrt{\frac{M}{\sigma^2}} (\bar{X}_M - \bar{x}) \leftarrow$$
 random variable of error

\Downarrow

Y_M converges weakly to Gaussian random variable

(mean: 0, variance: 1) as $M \rightarrow \infty$



④ Confidence interval

For quadrature rules, expectations converge as number of quadrature point goes to ∞ .

\Downarrow For empirical means

Under repeated sampling of empirical mean \bar{x}_M ,

a confidence interval is constructed.

True mean \bar{x} is contained in the interval

with some chosen probability P (Usually $P=0.95$)

⑤ Chebyshev's inequality

Given random variable \bar{X}_M (mean: \bar{x} , variable: σ^2), then Chebyshev's inequality states that

$$P(|Z_M| \geq k\sigma_M^2) \leq \frac{1}{k^2}, \quad k > 0$$

$$(Z_M = \bar{X}_M - \bar{x}, \sigma_M^2 = \frac{\sigma^2}{M})$$

proof

$$\chi_I \stackrel{\text{def}}{=} \begin{cases} 1 & \text{if } x \in I \quad (I \subset \mathbb{R}) \\ 0 & \text{otherwise} \end{cases}$$

And choose $I = \{x \in \mathbb{R} : |x| \geq k\sigma_M^2\}$. Then,

$$P(|Z_M| \geq k\sigma_M^2) = \mathbb{E}[\chi_I(Z_M)]$$

$$\leq \mathbb{E}\left[\left(\frac{Z_M}{k\sigma_M^2}\right)^2\right]$$

$$= \frac{1}{(k\sigma_M^2)^2} \mathbb{E}[Z_M^2] = \frac{1}{k^2} \sigma$$

CLT

• Strong law of large number

$$P\left(\lim_{n \rightarrow \infty} \bar{X}_n = \bar{x}\right) = 1 \quad (* \mathbb{E}[|X|] < \infty)$$

• Weak law of large number

$$\forall \varepsilon > 0, \lim_{n \rightarrow \infty} P(|\bar{X}_n - \bar{x}| > \varepsilon) = 0$$

Both laws imply that Monte Carlo approximations

\bar{X}_M converge to the true mean \bar{x} as $M \rightarrow \infty$

However, they don't imply how \bar{X}_M converge to \bar{x} . (\leftrightarrow Chebyshev's inequality do)

⑥ Monte Carlo approximation

For given PDF π_X , the Monte Carlo approximation to $\mathbb{E}[f(x)]$ is given by

$$\bar{f}_M = \sum_{i=1}^M w_i f(x_i) \quad \dots (*)$$

$$\left(\begin{matrix} * \\ * \end{matrix} \right. \left. \begin{matrix} \begin{pmatrix} x_1 \\ x_2 \\ \vdots \\ x_M \end{pmatrix} = \begin{pmatrix} X_1(w) \\ X_2(w) \\ \vdots \\ X_M(w) \end{pmatrix} \text{ is random samples.} \\ \text{The case of } w_i = \frac{1}{M} \text{ shows uniform weights.} \end{matrix} \right)$$

⑦ Generalised Monte Carlo approximation

This approximation to $\mathbb{E}[f(x)]$ use (*), but ...

(i) Use a different (and probably non-independent) joint distribution with PDF $\tilde{\pi}_X(x_1, x_2, \dots, x_M)$ for (X_n)

(ii) Use non-uniform weight $\{w_i\}$

$$\left(\begin{matrix} * \\ * \end{matrix} \right. \left. \begin{matrix} w_i > 0 \text{ and } \sum_{i=1}^M w_i = 1 \end{matrix} \right)$$

Then how to approximate by standard quadrature rules?

\rightarrow (a) select PDF as product of Dirac delta functions

$$\tilde{\pi}_M(x_1, x_2, \dots, x_M) = \prod_{i=1}^M \delta(x_i - c_i)$$

(b) properly written as ...

$$\text{measure: } \tilde{\mu}_M(dx_1, dx_2, \dots, dx_M) = \prod_{i=1}^M \mu_{c_i}(dx_i)$$

$$\text{weights: } w_i = \mu_{c_i}$$

However there is a difference between (a) and (b)

$$\rightarrow \text{For given quadrature rule } \begin{pmatrix} x_1 \\ x_2 \\ \vdots \\ x_M \end{pmatrix} = \begin{pmatrix} c_1 \\ c_2 \\ \vdots \\ c_M \end{pmatrix}$$

are always same. But in (b), the value of

$$\mu_M(dx) = \sum_{i=1}^M w_i \mu_{x_i}(dx)$$

will be different each time we draw from (X_1, X_2, \dots, X_M) with $\tilde{\pi}_X$

This make the case of $M \rightarrow \infty$ complicated.